Pedigree-based mixed linear models

The Prediction Problem

Model Equation

$$y = Xb + Zu + e$$

Other aspects of the model

First moments E[u] = 0, E[e] = 0, therefore E[y] = Xb

Second moments var[u] = G, var[e] = R, cov[u,e'] = 0

Distributional Assumptions e.g. u, e ~ MVN

Want to predict ${\bf u}$ or linear functions like ${\bf k'u}$

Original Solution

Generalized Least Squares (GLS)

For estimable $\mathbf{q}'\mathbf{b}$, $\mathbf{q}'\hat{\mathbf{b}}^0$ is BLUE (Best Linear Unbiased Estimator)

where
$$\hat{\mathbf{b}}^0 = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^T\mathbf{X}'\mathbf{V}^{-1}\mathbf{y}$$
 for $\mathbf{V} = \mathbf{Z}\mathbf{G}\mathbf{Z}' + \mathbf{R}$

then
$$\hat{\mathbf{u}} = \mathbf{G}\mathbf{Z}'\mathbf{V}^{-1}(\mathbf{y} - \mathbf{X}\hat{\mathbf{b}}^{0})$$
, is BLUP (BLU Predictor)

(same as Selection Index/BLP except $(y - X\hat{b}^0)$ in place of (y - Xb) obtained by exploiting (genetic) covariances between animals In traditional animal breeding practice

G is large and dense and determined by **A** the numerator relp matrix **V** is too big to compute **X'V**⁻¹

BLP vs GLS BLUP

$$y = X\beta + Zu + e$$

 $\mathbf{y} - \mathbf{X}\boldsymbol{\beta} = \mathbf{Z}\mathbf{u} + \mathbf{e}$, a fully random model

Selection Index Equations Pb = Gv

 $\mathbf{b} = \mathbf{P}^{-1}\mathbf{G}\mathbf{v}$, defines the best linear function to predict \mathbf{u} the "weights" are the same for every animal with the same sources of information (ie same traits observed)

BLP
$$\hat{\mathbf{u}} = \mathbf{b}'(\mathbf{y} - \mathbf{X}\boldsymbol{\beta}) = \mathbf{v}\mathbf{G}\mathbf{P}^{-1}(\mathbf{y} - \mathbf{X}\boldsymbol{\beta})$$

cf GLS BLUP
$$\hat{\mathbf{u}} = \mathbf{GZ'V^{-1}}(\mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}}^0)$$

Henderson's Contributions One

Developed methods to compute **G** and **R** from field data Henderson's Method I (not his!), II and III Including circumstances that involved selection

Henderson's Contributions Two

Invented the Mixed Model Equations

$$\begin{bmatrix} \mathbf{X}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{X}'\mathbf{R}^{-1}\mathbf{Z} \\ \mathbf{Z}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{Z}'\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1} \end{bmatrix} \begin{bmatrix} \hat{\mathbf{b}}^{0} \\ \hat{\mathbf{u}} \end{bmatrix} = \begin{bmatrix} \mathbf{X}'\mathbf{R}^{-1}\mathbf{y} \\ \mathbf{Z}'\mathbf{R}^{-1}\mathbf{y} \end{bmatrix}, \text{ for full rank } \mathbf{G}$$

and jointly showed $k'\hat{b}^0$ and \hat{u} were BLUE and BLUP

Computationally tractable if **G** and **R** assumed diagonal or block-diagonal (eg sire model with relationships ignored)

(Order 40 matrix takes weeks to invert by hand)

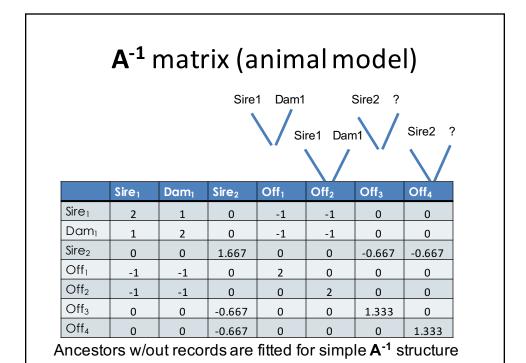
MME typically sparse in national animal evaluation

Example NRM or **A** matrix



	Offspring ₁	Offspring ₂	Offspring ₃	Offspring ₄
Offspring ₁	1	1/2	0	0
Offspring ₂	1/2	1	0	0
Offspring ₃	0	0	1	1/4
Offspring ₄	0	0	1/4	1

Sires and dams unrelated and non-inbred Simple calculation of \mathbf{A}_{p}^{-1} requires including all ancestors and would result in a matrix of order 7 not 4



Henderson's Contributions Three

Invented an algorithm to directly form A^{-1} from a pedigree list Then G^{-1} can be formed as a scalar product or kronecker product define d to be "mendelian" sampling variance d = (1, 3/4, 1/2) for 0, 1 or 2 parents known define s' = (-1/2, -1/2, 1) to represent sire (if known), dam (if known

define $\mathbf{s}' = (-1/2, -1/2, 1)$ to represent sire (if known), dam (if known) and individual equations

accumulate $\mathbf{s}d^{-1}\mathbf{s}'$ in the sire, dam and individual rows/columns for every trio of animals in the pedigree list

Consequence of A⁻¹ structure

sira dam i

Accumulate for each animal

sire
$$\begin{bmatrix} 0.25 & 0.25 & -0.5 \\ dam & 0.25 & 0.25 & -0.5 \\ -0.5 & -0.5 & 1 \end{bmatrix} d^{-1}$$

When both parents are known

Nonparents (ie terminal offspring)

Own equation (ie row) has 2 on diagonal, -1 in sire column -1 in dam column Parent with one offspring

Own equation has 2+1/2 on diagonal, -1 in sire and dam columns in addition to -1/2 in the column of its mate, -1 in column of offspring

Parent with many offspring to different mates accumulates a large diagonal element, many small negative offdiagonals

Consider rearranging the MME

In general,

$$\begin{bmatrix} \mathbf{Z}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{Z}'\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1} \end{bmatrix} \begin{bmatrix} \hat{\mathbf{b}}^{0} \\ \hat{\mathbf{u}} \end{bmatrix} = \begin{bmatrix} \mathbf{Z}'\mathbf{R}^{-1}\mathbf{y} \end{bmatrix}$$
or equivalently $\begin{bmatrix} \mathbf{Z}'\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1} \end{bmatrix} [\hat{\mathbf{u}}] = \begin{bmatrix} \mathbf{Z}'\mathbf{R}^{-1} (\mathbf{y} - \mathbf{X}\hat{\mathbf{b}}^{0}) \end{bmatrix}$
Single trait animal model $\mathbf{R} = \mathbf{I}\sigma_{e}^{2}$, $\mathbf{G} = \mathbf{A}\sigma_{g}^{2}$, $\mathbf{G}^{-1} = \mathbf{A}^{-1}\sigma_{g}^{-2}$
or multiplying σ_{e}^{2} , $\begin{bmatrix} \mathbf{Z}'\mathbf{Z} + \lambda \mathbf{A}^{-1} \end{bmatrix} [\hat{\mathbf{u}}] = \begin{bmatrix} \mathbf{Z}'(\mathbf{y} - \mathbf{X}\hat{\mathbf{b}}^{0}) \end{bmatrix}$, with $\lambda = \frac{\sigma_{e}^{2}}{\sigma_{g}^{2}}$

Consider the MME for a nonparent

$$\begin{aligned} & \left[\mathbf{Z}'\mathbf{Z} + \lambda \mathbf{A}^{-1} \right] \left[\hat{\mathbf{u}} \right] = \left[\mathbf{Z}' \left(\mathbf{y} - \mathbf{X} \hat{\mathbf{b}}^{0} \right) \right] \\ & \text{Nonparent animal with one record} \\ & (1 + 2\lambda) \hat{u}_{animal} - \lambda \hat{u}_{sire} - \lambda \hat{u}_{dam} = adjusted _y \\ & \hat{u}_{animal} = \frac{2\lambda \left(\hat{u}_{sire} + \hat{u}_{dam} \right)}{(1 + 2\lambda)2} + \frac{\left(adjusted _y \right)}{(1 + 2\lambda)} \\ & = (1 - w)PA + w \left(adjusted _y \right) \quad for \quad w = \frac{1}{(1 + 2\lambda)} \end{aligned}$$

Consider the MME for a nonparent

$$\hat{u}_{animal} = (1 - w)PA + w(adjusted_y)$$
 for $w = \frac{1}{(1 + 2\lambda)}$

$$\lambda = \frac{1 - h^2}{h^2} so for h^2 = 1, \ \lambda = 0, w = 1, \ (no \ shrinkage)$$

for
$$h^2 = low$$
, $\lambda = big$, $w = small$, (shrink the deviation)

Two sources of BV information are pooled

The parent average PA

The individual prediction (shrunk deviation) with heritability influencing shrinkage

Consider the MME for a nonparent

$$\left[\mathbf{Z}'\mathbf{Z} + \lambda \mathbf{A}^{-1}\right] \left[\hat{\mathbf{u}}\right] = \left[\mathbf{Z}'\left(\mathbf{y} - \mathbf{X}\hat{\mathbf{b}}^{0}\right)\right]$$

Nonparent animal with one record

$$\hat{u}_{animal} = (1 - w)PA + w(adjusted y)$$

Nonparent animal with no record

$$2\lambda\hat{u}_{animal} - \lambda\hat{u}_{sire} - \lambda\hat{u}_{dam} = 0$$

$$\hat{u}_{animal} = \frac{\lambda (\hat{u}_{sire} + \hat{u}_{dam})}{\lambda 2} = \frac{(\hat{u}_{sire} + \hat{u}_{dam})}{2} = PA$$

Reliability of nonparents

Property of BLP/BLUP is $cov(u, \hat{u}) = var(\hat{u})$ so $r^2 = \frac{var(\hat{u})}{var(u)}$

but $\hat{u}_{nonparent} = \frac{\hat{u}_{sire}}{2} + \frac{\hat{u}_{dam}}{2}$, for nonparent without a record

so
$$\mathbf{r}_{nonparent}^2 = \frac{r_{sire}^2}{4} + \frac{r_{dam}^2}{4} \le \frac{1}{2}$$

Finally $\Delta G = \frac{ir_{nonparent}\sigma_g}{L}$, limiting selection response

when candidates at puberty lack phenotypic information

An option to do better

Solution

- We need a different representation of the covariance between relatives, that allows relatives other than parents to directly contribute to the prediction of nonparents without records
- The NRM or A-matrix is an expectation of relationships in the context of repeated sampling of the pedigree (conditional on pedigree)

A-matrix

- Relationship with self is 1+F (noninbred F=0)
- (Additive) relationship of ½ between non-inbred full-sibs and between parents and non-inbred offspring
- Relationship of ¼ between non-inbred half-sibs and between grandparents and offspring
- But particular individuals can have greater or lesser values
 - If we know their genotype we can compute relationships conditional on the chromosome regions they inherited

Relationship matrix A matrix $\begin{bmatrix} 1 & 0 & .5 & .5 & .5 & .5 & .5 \\ 0 & 1 & .5 & .5 & .5 & .5 & .5 \\ .5 & .5 & 1 & .5 & .5 & .5 & .5 \\ .5 & .5 & .5 & .5 & .5 & .5 & .5 \\ .5 & .5 & .5 & .5 & .5 & .5 & .5 \\ .5 & .5 & .5 & .5 & .5 & .5 & .1 \end{bmatrix}$ Consider a sire, dam and 4 full sibs A-inverse matrix $\begin{bmatrix} 3 & 2 & -1 & -1 & -1 & -1 \\ 2 & 3 & -1 & -1 & -1 & -1 \\ -1 & -1 & 2 & 0 & 0 & 0 \\ -1 & -1 & 0 & 2 & 0 & 0 \\ -1 & -1 & 0 & 0 & 2 & 0 \\ -1 & -1 & 0 & 0 & 0 & 2 \end{bmatrix}$

Relationship matrix **G** matrix A matrix .5 .5 .5 .5 1 .5 .5 .5 .5 .5 .5 .5 1 .5 .5 .5 .5 .5 .5 .5 .5 1 .5 G-inverse matrix A-inverse matrix 3.5 2.5 -1.25-1.25-1.25-1.252.5 3.5 -1.25-1.25-1.25-1.253 -1 -1 -1 -1 -1.25 -1.252.1875 0.3125 -0.31250.3125 -1 -1 2 0 0 0 -1.25 -1.25 -0.31252.1875 0.3125 0.3125 -1 -1 0 2 0 0 -1.25 -1.250.3125 -0.3125 0.3125 2.1875 -1.25-1.250.3125 0.3125 -0.3125 2.1875

Predict the last animal with no data

$$\begin{bmatrix} -1.25\hat{u}_{sire} & -1.25\hat{u}_{dam} & .3125\hat{u}_{sib1} & .3125\hat{u}_{sib2} & -.3125\hat{u}_{sib3} & 2.1875\hat{u}_{candidate} \end{bmatrix} = \begin{bmatrix} 0 \end{bmatrix}$$

$$\hat{u}_{candidate} = \frac{1.25(\hat{u}_{sire} + \hat{u}_{dam}) - 0.3125(\hat{u}_{sib1} + \hat{u}_{sib2}) + 0.3125\hat{u}_{sib3}}{2.1875}$$

But to form **G**, we needed to know which loci/QTL contribute to variation in performance

Some MME Results

$$\begin{bmatrix} X'X & X'Z \\ Z'X & Z'Z + \lambda G^{-1} \end{bmatrix}^{-} = \begin{bmatrix} C^{11} & C^{12} \\ C^{21} & C^{22} \end{bmatrix}$$

$$var(\mathbf{g}) = \mathbf{G} \quad var(\hat{\mathbf{g}}) = \mathbf{G} \cdot \mathbf{C}^{22} \quad var(\hat{\mathbf{g}} \cdot \mathbf{g}) = \mathbf{C}^{22} \quad r_{g\hat{g}}^{2} = \frac{var(\hat{g})}{var(g)}$$
$$var(\mathbf{k}'\mathbf{g}) = \mathbf{k}'\mathbf{G}\mathbf{k} \quad var(\mathbf{k}'\hat{\mathbf{g}}) = \mathbf{k}'\left(\mathbf{G} \cdot \mathbf{C}^{22}\right)\mathbf{k}$$