# Pedigree-based mixed linear models

### The Prediction Problem

Model Equation

$$y = Xb + Zu + e$$

Other aspects of the model

First moments E[u] = 0, E[e] = 0, therefore E[y] = Xb

Second moments var[u] = G, var[e] = R, cov[u,e'] = 0

Distributional Assumptions e.g. **u**, **e** ~ MVN

Want to predict u or linear functions like k'u

## **Original Solution**

Generalized Least Squares (GLS)

For estimable  $\mathbf{q'b}$ ,  $\mathbf{q'\hat{b}^0}$  is BLUE (Best Linear Unbiased Estimator)

where 
$$\hat{\mathbf{b}}^0 = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^T\mathbf{X}'\mathbf{V}^{-1}\mathbf{y}$$
 for  $\mathbf{V} = \mathbf{Z}\mathbf{G}\mathbf{Z}' + \mathbf{R}$ 

then  $\hat{\mathbf{u}} = \mathbf{G}\mathbf{Z}'\mathbf{V}^{-1}(\mathbf{y} - \mathbf{X}\hat{\mathbf{b}}^{0})$ , is BLUP (BLU Predictor)

(same as Selection Index/BLP except  $(\mathbf{y} - \mathbf{X}\hat{\mathbf{b}}^0)$  in place of  $(\mathbf{y} - \mathbf{X}\mathbf{b})$ 

obtained by exploiting (genetic) covariances between animals

In traditional animal breeding practice

**G** is large and dense and determined by **A** the numerator relp matrix

V is too big to compute  $X'V^{-1}$ 

#### **BLP vs GLS BLUP**

$$y = X\beta + Zu + e$$
  
 $y - X\beta = Zu + e$ , a fully random model  
Selection Index Equations  $Pb = Gv$ 

 $\mathbf{b} = \mathbf{P}^{-1}\mathbf{G}\mathbf{v}$ , defines the best linear function to predict  $\mathbf{u}$  the "weights" are the same for every animal with the same sources of information (ie same traits observed)

BLP 
$$\hat{\mathbf{u}} = \mathbf{b}'(\mathbf{y} - \mathbf{X}\boldsymbol{\beta}) = \mathbf{v}\mathbf{G}\mathbf{P}^{-1}(\mathbf{y} - \mathbf{X}\boldsymbol{\beta})$$

cf GLS BLUP  $\hat{\mathbf{u}} = \mathbf{G}\mathbf{Z}'\mathbf{V}^{-1}(\mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}}^{0})$ 

#### Henderson's Contributions One

Developed methods to compute **G** and **R** from field data Henderson's Method I (not his!), II and III Including circumstances that involved selection

### Henderson's Contributions Two

Invented the Mixed Model Equations

$$\begin{bmatrix} \mathbf{X'R^{-1}X} & \mathbf{X'R^{-1}Z} \\ \mathbf{Z'R^{-1}X} & \mathbf{Z'R^{-1}Z + G^{-1}} \end{bmatrix} \begin{bmatrix} \hat{\mathbf{b}}^{0} \\ \hat{\mathbf{u}} \end{bmatrix} = \begin{bmatrix} \mathbf{X'R^{-1}y} \\ \mathbf{Z'R^{-1}y} \end{bmatrix}, \text{ for full rank } \mathbf{G}$$

and jointly showed  $\mathbf{k'}\hat{\mathbf{b}}^0$  and  $\hat{\mathbf{u}}$  were BLUE and BLUP

Computationally tractable if G and R assumed diagonal or block-diagonal (eg sire model with relationships ignored)

(Order 40 matrix takes weeks to invert by hand)

MME typically sparse in national animal evaluation

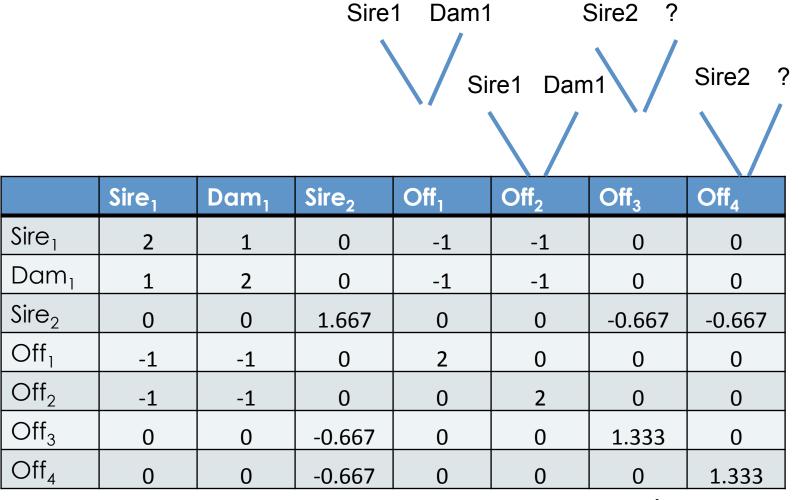
### Example NRM or **A** matrix



	Offspring <sub>1</sub>	Offspring <sub>2</sub>	Offspring <sub>3</sub>	Offspring <sub>4</sub>
Offspring <sub>1</sub>	1	1/2	0	0
Offspring <sub>2</sub>	1/2	1	0	0
Offspring <sub>3</sub>	0	0	1	1/4
Offspring <sub>4</sub>	0	0	1/4	1

Sires and dams unrelated and non-inbred Simple calculation of  $A_p^{-1}$  requires including all ancestors and would result in a matrix of order 7 not 4

# A<sup>-1</sup> matrix (animal model)



Ancestors w/out records are fitted for simple A<sup>-1</sup> structure

### Henderson's Contributions Three

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Invented an algorithm to directly form A^{-1} from a pedigree list Then G^{-1} can be formed as a scalar product or kronecker product define d to be "mendelian" sampling variance d = (1, 3/4, 1/2) for 0, 1 or 2 parents known define \mathbf{s'} = (-1/2, -1/2, 1) to represent sire (if known), dam (if known) and individual equations accumulate \mathbf{s}d^{-1}\mathbf{s'} in the sire, dam and individual rows/columns
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for every trio of animals in the pedigree list

# Consequence of A<sup>-1</sup> structure

When both parents are known

Nonparents (ie terminal offspring)

Own equation (ie row) has 2 on diagonal, -1 in sire column -1 in dam column Parent with one offspring

Own equation has 2+1/2 on diagonal, -1 in sire and dam columns

in addition to -1/2 in the column of its mate, -1 in column of offspring

Parent with many offspring to different mates accumulates a large diagonal element, many small negative offdiagonals

# Consider rearranging the MME

In general,

$$\begin{bmatrix} \mathbf{Z}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{Z}'\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1} \end{bmatrix} \begin{bmatrix} \hat{\mathbf{b}}^{0} \\ \hat{\mathbf{u}} \end{bmatrix} = \begin{bmatrix} \mathbf{Z}'\mathbf{R}^{-1}\mathbf{y} \end{bmatrix}$$
or equivalently  $\begin{bmatrix} \mathbf{Z}'\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1} \end{bmatrix} [\hat{\mathbf{u}}] = \begin{bmatrix} \mathbf{Z}'\mathbf{R}^{-1} (\mathbf{y} - \mathbf{X}\hat{\mathbf{b}}^{0}) \end{bmatrix}$ 
Single trait animal model  $\mathbf{R} = \mathbf{I}\sigma_{e}^{2}$ ,  $\mathbf{G} = \mathbf{A}\sigma_{g}^{2}$ ,  $\mathbf{G}^{-1} = \mathbf{A}^{-1}\sigma_{g}^{-2}$ 
or multiplying  $\sigma_{e}^{2}$ ,  $\begin{bmatrix} \mathbf{Z}'\mathbf{Z} + \lambda \mathbf{A}^{-1} \end{bmatrix} [\hat{\mathbf{u}}] = \begin{bmatrix} \mathbf{Z}'(\mathbf{y} - \mathbf{X}\hat{\mathbf{b}}^{0}) \end{bmatrix}$ , with  $\lambda = \frac{\sigma_{e}^{2}}{\sigma_{g}^{2}}$ 

# Consider the MME for a nonparent

$$\left[ \mathbf{Z'Z} + \lambda \mathbf{A^{-1}} \right] \left[ \hat{\mathbf{u}} \right] = \left[ \mathbf{Z'} \left( \mathbf{y} - \mathbf{X} \hat{\mathbf{b}}^{0} \right) \right]$$

Nonparent animal with one record

$$(1+2\lambda)\hat{u}_{animal} - \lambda\hat{u}_{sire} - \lambda\hat{u}_{dam} = adjusted \_y$$

$$\hat{u}_{animal} = \frac{2\lambda(\hat{u}_{sire} + \hat{u}_{dam})}{(1+2\lambda)2} + \frac{(adjusted \_y)}{(1+2\lambda)}$$

$$= (1-w)PA + w(adjusted \_y) \text{ for } w = \frac{1}{(1+2\lambda)}$$

# Consider the MME for a nonparent

$$\hat{u}_{animal} = (1 - w)PA + w(adjusted y)$$
 for  $w = \frac{1}{(1 + 2\lambda)}$ 

$$\lambda = \frac{1 - h^2}{h^2} so for h^2 = 1, \ \lambda = 0, w = 1, \ (no \ shrinkage)$$

for  $h^2 = low$ ,  $\lambda = big$ , w = small, (shrink the deviation)

Two sources of BV information are pooled

The parent average PA

The individual prediction (shrunk deviation)

with heritability influencing shrinkage

# Consider the MME for a nonparent

$$\left[\mathbf{Z'Z} + \lambda \mathbf{A^{-1}}\right] \left[\hat{\mathbf{u}}\right] = \left[\mathbf{Z'}\left(\mathbf{y} - \mathbf{X}\hat{\mathbf{b}}^{0}\right)\right]$$

Nonparent animal with one record

$$\hat{u}_{animal} = (1 - w)PA + w(adjusted y)$$

Nonparent animal with no record

$$2\lambda\hat{u}_{animal} - \lambda\hat{u}_{sire} - \lambda\hat{u}_{dam} = 0$$

$$\hat{u}_{animal} = \frac{\lambda(\hat{u}_{sire} + \hat{u}_{dam})}{\lambda 2} = \frac{(\hat{u}_{sire} + \hat{u}_{dam})}{2} = PA$$

# Reliability of nonparents

Property of BLP/BLUP is  $cov(u, \hat{u}) = var(\hat{u})$  so  $r^2 = \frac{var(u)}{var(u)}$ 

but  $\hat{u}_{nonparent} = \frac{\hat{u}_{sire}}{2} + \frac{\hat{u}_{dam}}{2}$ , for nonparent without a record

so 
$$r_{nonparent}^2 = \frac{r_{sire}^2}{4} + \frac{r_{dam}^2}{4} \le \frac{1}{2}$$

Finally  $\Delta G = \frac{ir_{nonparent}\sigma_g}{L}$ , limiting selection response

when candidates at puberty lack phenotypic information

# An option to do better

### Solution

- We need a different representation of the covariance between relatives, that allows relatives other than parents to directly contribute to the prediction of nonparents without records
- The NRM or A-matrix is an expectation of relationships in the context of repeated sampling of the pedigree (conditional on pedigree)

#### **A**-matrix

- Relationship with self is 1+F (noninbred F=0)
- (Additive) relationship of ½ between non-inbred full-sibs and between parents and non-inbred offspring
- Relationship of ¼ between non-inbred half-sibs and between grandparents and offspring
- But particular individuals can have greater or lesser values
  - If we know their genotype we can compute relationships conditional on the chromosome regions they inherited

## Relationship matrix

#### **A** matrix

 1
 0
 .5
 .5
 .5
 .5

 0
 1
 .5
 .5
 .5
 .5

 .5
 .5
 1
 .5
 .5
 .5

 .5
 .5
 .5
 .5
 .5
 .5

 .5
 .5
 .5
 .5
 .5
 .1

 .5
 .5
 .5
 .5
 .5
 .1

#### **A-inverse** matrix

$$\begin{bmatrix} 3 & 2 & -1 & -1 & -1 & -1 \\ 2 & 3 & -1 & -1 & -1 & -1 \\ -1 & -1 & 2 & 0 & 0 & 0 \\ -1 & -1 & 0 & 2 & 0 & 0 \\ -1 & -1 & 0 & 0 & 2 & 0 \\ -1 & -1 & 0 & 0 & 0 & 2 \end{bmatrix}$$

Consider a sire, dam and 4 full sibs

### Relationship matrix

#### **A** matrix

1	0	.5	.5	.5	.5
0	1	.5	.5	.5	.5
.5	.5	1	.5	.5	.5
.5	.5	.5	1	.5	.5
.5	.5	.5	.5	1	.5
.5	.5	.5	.5	.5	1

#### **G** matrix

1	0	.5	.5	.5	.5
0	1	.5	.5	.5	
.5	.5	1	.6	.4	.4
.5	.5	.6	1	.4	.4
.5	.5	.4	.4	1	.6
.5	.5	.4	.4	.6	1

$$\begin{bmatrix} 3 & 2 & -1 & -1 & -1 & -1 \\ 2 & 3 & -1 & -1 & -1 & -1 \\ -1 & -1 & 2 & 0 & 0 & 0 \\ -1 & -1 & 0 & 2 & 0 & 0 \\ -1 & -1 & 0 & 0 & 2 & 0 \\ -1 & -1 & 0 & 0 & 0 & 2 \end{bmatrix}$$

#### **G-inverse** matrix

#### Predict the last animal with no data

$$\begin{bmatrix} -1.25\hat{u}_{sire} & -1.25\hat{u}_{dam} & .3125\hat{u}_{sib1} & .3125\hat{u}_{sib2} & -.3125\hat{u}_{sib3} & 2.1875\hat{u}_{candidate} \end{bmatrix} = \begin{bmatrix} 0 \end{bmatrix}$$

$$\hat{u}_{candidate} = \frac{1.25 \left(\hat{u}_{sire} + \hat{u}_{dam}\right) - 0.3125 \left(\hat{u}_{sib1} + \hat{u}_{sib2}\right) + 0.3125 \hat{u}_{sib3}}{2.1875}$$

But to form **G**, we needed to know which loci/QTL contribute to variation in performance

### Some MME Results

$$\begin{bmatrix} X'X & X'Z \\ Z'X & Z'Z + \lambda G^{-1} \end{bmatrix}^{-} = \begin{bmatrix} C^{11} & C^{12} \\ C^{21} & C^{22} \end{bmatrix}$$

$$var(\mathbf{g}) = \mathbf{G} \quad var(\hat{\mathbf{g}}) = \mathbf{G} \cdot \mathbf{C}^{22} \quad var(\hat{\mathbf{g}} - \mathbf{g}) = \mathbf{C}^{22} \quad r_{g\hat{g}}^{2} = \frac{var(\hat{g})}{var(g)}$$
$$var(\mathbf{k}'\mathbf{g}) = \mathbf{k}'\mathbf{G}\mathbf{k} \quad var(\mathbf{k}'\hat{\mathbf{g}}) = \mathbf{k}'\left(\mathbf{G} \cdot \mathbf{C}^{22}\right)\mathbf{k}$$