An Equivalent (animal) Model for Genomic Prediction

Usual Single Trait Pedigree Model

$$y = Xb + Zu + e$$

$$\begin{bmatrix} X'X & X'Z \\ Z'X & Z'Z + \lambda A^{-1} \end{bmatrix} \begin{bmatrix} b^0 \\ \widehat{u} \end{bmatrix} \begin{bmatrix} X'y \\ Z'y \end{bmatrix}$$

Single Trait Marker Effects Model

$$y = Xb + ZM\alpha + e$$

$$egin{bmatrix} X'X & X'ZM \ M'Z'X & M'Z'ZM + rac{\sigma_e^2}{\sigma_lpha^2}I \end{bmatrix} egin{bmatrix} b^0 \ \widehat{lpha} \end{bmatrix} egin{bmatrix} X'y \ M'Z'y \end{bmatrix}$$

$$egin{bmatrix} X'X & X'M \ M'X & M'M + rac{\sigma_e^2}{\sigma_a^2} I \end{bmatrix} egin{bmatrix} b^0 \ \widehat{lpha} \end{bmatrix} egin{bmatrix} X'y \ M'y \end{bmatrix} \ Then \ \widehat{u} = M \widehat{lpha} \end{aligned}$$

This is MHG "BLUP" or is sometimes known as (Ridge-regression) RR-BLUP with known $\frac{\sigma_e^2}{\sigma_\sigma^2}$

More loci than animals

But for selection we are more interested in animal (not allelic) merit

$$y = Xb + ZM\alpha + e$$

$$y = Xb + M\alpha + e$$

$$y = Xb + IM\alpha + e$$

$$y = Xb + I'Z'''u'' + e$$

Order of MME is number of fixed effects plus number of animals Consider the implications for 100-1,000 animals with 50,000 loci

Mixed Model Equations

$$y = Xb + IM\alpha + e$$

$$y = Xb + "Z" "u" + e$$

$$egin{bmatrix} X'X & X' \ X & I + \sigma_e^2 [var(Mlpha)]^{-1} \end{bmatrix} egin{bmatrix} b^0 \ \widehat{Mlpha} \end{bmatrix} X'y \ y \end{bmatrix}$$

What is $var(M\alpha)$?

$$var(M\alpha) = Mvar(\alpha)M'$$

= $MIM'\sigma_{\alpha}^{2} = MM'\sigma_{\alpha}^{2}$

Homogenous locus variance

What is $var(M\alpha)$?

$$var(M\alpha) = Mvar(\alpha)M'$$

= $MIM'\sigma_{\alpha}^{2} = MM'\sigma_{\alpha}^{2}$

Homogeneous locus variance

$$egin{aligned} var[u] &= var[\sum^{loci} m_i lpha_i] = \sum^{loci} var(m_i lpha_i) \ &= \sum^{loci} m_i var(lpha_i) \, m_i' = \sum^{loci} m_i m_i' \sigma_{lpha_i}^2 \end{aligned}$$

Heterogeneous locus variance

Genomic Relationship Matrix

$$G = MM'$$

$$G=\sum^{loci}m_im_i^{'}\sigma_{lpha_i}^2$$

Trait specific

Genomic Relationship Matrix

$$M = k \text{ columns of } (0, 1, 2) \text{ marker covariates}$$

$$G = [MM' + (2 - M)(2 - M)']/k$$

$$var[u] = G\sigma_a^2$$

GBLUP

$$egin{aligned} \left[egin{aligned} X'X & X'Z & \left[egin{aligned} b' \ \widehat{u} \end{array}
ight] \left[egin{aligned} X'y \ Z'y \end{aligned}
ight] \ G &= MM' \ \widehat{u} &= M\widehat{lpha} \end{aligned}$$

GBLUP

- If the variance parameters are assumed known and the inverse of the genomic relationship matrix is multiplied by (known) λ , the system is known as GBLUP, as opposed to conventional pedigree or PBLUP
 - It is effectively weighting all the loci equally
 - It is similar to BayesC0 except that in that method we estimate the variance components after including a prior distribution for them

Lack of Equivalence

- The GBLUP and Marker Effects Models (MEM) such as BayesCO with high df for the prior variances will give the same EBV for the genotyped animals
 - This is true regardless of
 - whether the models fit the A allele at every locus, the B allele at every locus, or both alleles at every locus
 - how the alleles are centered (coded 0,1,2 or -1,0,1 etc)
 - However, the PEV (and reliability) for GBLUP are not invariant to these alternative models

Genomic Analysis Combining Genotyped and Non-Genotyped Individuals

Why a Combined Analysis?

- To exploit all the available phenotypic data in GWAS and genomic prediction
 - Not just the records on genotyped individuals
 - Account for preselection of genotyped individuals
- To ensure that genomic predictions include all available information
- To avoid approximations required in multistep analyses (that lead to double-counting)

Multi-step Genomic Prediction Analysis

- Mixed model evaluation using all phenotypes and pedigree information to generate EBV and R²
- Deregression of EBV on genotyped individuals using EBV and R² of trios of every genotyped individual, its sire and its dam
- Weighted multiple regression analysis of deregressed EBV to estimate SNP effects
- Genomic prediction DGV of genotyped individuals
- Pedigree prediction of DGV for nongenotyped
- Selection Index blending of DGV & EBV for GE-EBV

Pedigree Prediction

$$\begin{bmatrix} y_n \\ y_g \end{bmatrix} = \begin{bmatrix} X_n \\ X_g \end{bmatrix} b + \begin{bmatrix} Z_n & 0 \\ 0 & Z_g \end{bmatrix} \begin{bmatrix} u_n \\ u_g \end{bmatrix} + \begin{bmatrix} e_n \\ e_g \end{bmatrix}$$

with

$$varig[egin{array}{c} u_n \ u_g \end{array} = egin{bmatrix} A_{nn} & A_{ng} \ A_{gn} & A_{gg} \end{array} \sigma_a^2$$

Where **A** is the numerator relationship matrix (from pedigree) with subscripts n=non-genotyped & g=genotyped

Nejati-Javaremi et al (1997)

Replace A with G

 $M = k \ columns \ of \ (0, 1, 2) \ marker \ covariates$

$$G = [MM' + (2 - M)(2 - M)']/k$$

Various other authors expanded this with various approaches to center the marker covariates to create a Genomic Relationship Matrix

Fitting G⁻¹ in the mixed model equations is known as GBLUP and gives the same estimates of genomic merit as MHG "BLUP"

Genotyped Animals

$$y_g = X_g b + Z_g u_g + e_g$$

Meuwissen, Hayes & Goddard (2001)

$$with \ u_g = M_g lpha = \sum_{j=1}^{j=\#loci} m_j lpha_j \delta_j$$

 $\alpha_j = substitution\ effect$

$$\delta_j = (0,1) indicator variable$$

Bayesian Alphabet

$$egin{aligned} eta_j &= 1, \;\; \sigma_{lpha_j}^2 = (known)\,\sigma_{lpha}^2\,was\,"BLUP" \ eta_j &= 1, \;\; \sigma_{lpha_j}^2 = (unknown)\,\sigma_{lpha_j}^2\,was\,BayesA \ egin{aligned} eta_j &= 0\,with\,known\,probability = \pi \ \sigma_{lpha_j}^2 &= (unknown)\,\sigma_{lpha_j}^2\,was\,BayesB \end{aligned}$$

Meuwissen, Hayes & Goddard (2001)

$$\delta_{j} = 0 \text{ with } (un) \text{ known probability} = \pi$$

$$\sigma_{\alpha_{j}}^{2} = (unknown) \sigma_{\alpha}^{2} \text{ was BayesC or } (BayesC\pi)$$

Kizilkaya et al (2010); Habier et al (2011)

Evolution of "The Model"

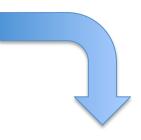
Genomic Relationship Matrix

$$y = Xb + Z\mathbf{u} + e$$

M = k columns of (0, 1, 2) marker covariatesG = [MM' + (2 - M)(2 - M)']/k

 $var[\mathbf{u}] = \mathbf{G}\sigma_a^2, var[e] = I\sigma_e^2$

Nejati-Javaremi et al. (1997)



Equivalent

Breeding Value Model

$$var[u] = var[Mlpha] = MIM'\sigma_lpha^2$$

Stranden & Garrick (2009)

Pedigree Relationship Matrix

$$y = Xb + Zu + e$$

 $var[u] = \mathbf{A}\sigma_a^2, var[e] = I\sigma_e^2$

Breeding Value Model

 $u = M\alpha = sum \ of \ substitution \ effects$

$$y = Xb + ZM\alpha + e$$

 $var[lpha] = I\sigma_{lpha}^{\scriptscriptstyle 2}, var[e] = I\sigma_{\scriptscriptstyle e}^{\scriptscriptstyle 2}$

Meuwissen et al. (2001)

Marker Effects Model (MEM)



What to do with the non-genotyped?

Known as Single-Step "First Attempt"

$$varegin{bmatrix} u_n \ u_g \end{bmatrix} = egin{bmatrix} A_{nn} & A_{ng} \ A_{gn} & G_{gg} \end{bmatrix} \sigma_a^2$$

Just replace that part of the numerator relationship matrix with genomic relationships

Then need a "brute-force" inversion of the var-cov matrix

What to do with the non-genotyped?

Known as Single-Step "Second Attempt" (with brute force inverse)

$$H = var igg[u_n igg] \sigma_a^{-2} = igg[A_{nn} + A_{ng} A_{gg}^{-1} G_{gg} A_{gg}^{-1} A_{gn} \quad A_{ng} A_{gg}^{-1} G_{gg} igg] \ G_{gg} A_{gg}^{-1} A_{gn} \quad G_{gg} igg]$$
 Legarra et al (2009)

Then with recognition of its simply structured inverse

$$H^{-1} = A^{-1} + egin{bmatrix} 0 & 0 \ 0 & G_{qq}^{-1} - A_{qq}^{-1} \end{bmatrix}$$

Aguilar et al (2010)

Offering programming appeal by simply replacing A⁻¹ in MME by H⁻¹ known as Single-Step GBLUP and variants of which are widely used

 Its predictive ability can be improved by introducing another ad hoc constant κ whose optimal value can be found by trial and error

$$H^{-1} = A^{-1} + egin{bmatrix} 0 & 0 \ 0 & arkappa(G_{gg}^{-1} - A_{gg}^{-1}) \end{bmatrix}$$

- When there are less loci than genotyped individuals, G is singular
- When there are more loci than genotyped individuals, G is singular if locus covariates are centered by allele frequency

(since G=MM' and M'1=0 then G1=0)

 These problems can be overcome by adhoc regression of G towards A

- The var-cov matrix involves a blending of A and G requiring that they represent the same "base"
 - The base in A is the pedigree founders but the allele frequencies are not usually known in that population
- It is not clear what to use to center locus covariates in populations of mixed breeds, or populations with variable breed percentages

Issues with single-step GBLUP

- The matrix G is often singular
 - More animals than markers
 - If G is centered with observed allele frequency
- The matrix G must be "on the same base" as A

Rather than using $G_{gg}^{-1} - A_{gg}^{-1}$

The model is tuned using

$$\tau[a+b((1-c)G_{gg}+cA_{gg})]^{-1}-\omega A_{gg}^{-1}$$

with some trial and error and

$$\tau \le 1$$
; $\omega \le 1$; $a \le 0.1$; $b \le 1$; $0.05 \le c \le 0.2$

Computing effort increases with numbers genotyped

- It requires brute force inversion of 2 matrices whose order is the number of genotyped individuals (ie **G** and A_{gg})
 - The inversion effort increase rapidly with number of genotyped individuals
 - Inversion is impractical beyond say 100,000 individuals
- Ignacy now has an approximation approach for computing these inverses

- It is not computationally straightforward for extension to Single-Step BayesA
- It is not suitable for application of mixture models (BayesB, BayesC, BayesCπ)
 - But these models that provide variable selection are particularly appealing in fine-mapping applications such as with imputed NGS genotypes

Let's revisit the basic idea

$$egin{aligned} egin{aligned} egi$$

Substituting these results gives

$$\begin{bmatrix} y_n \\ y_g \end{bmatrix} = \begin{bmatrix} X_n \\ X_g \end{bmatrix} b + \begin{bmatrix} Z_n & 0 \\ 0 & Z_g \end{bmatrix} \begin{bmatrix} u_n \\ u_g \end{bmatrix} + \begin{bmatrix} e_n \\ e_g \end{bmatrix}$$

$$=egin{bmatrix} X_n \ X_g \end{bmatrix} b + egin{bmatrix} Z_n & 0 \ 0 & Z_g \end{bmatrix} egin{bmatrix} A_{ng}A_{gg}^{-1}M_glpha \ M_glpha \end{bmatrix} + egin{bmatrix} Z_n & 0 \ 0 & 0 \end{bmatrix} egin{bmatrix} oldsymbol{arepsilon}_n \ 0 \end{bmatrix} + egin{bmatrix} e_n \ e_g \end{bmatrix}$$

$$= \begin{bmatrix} X_n \\ X_q \end{bmatrix} b + \begin{bmatrix} Z_n A_{ng} A_{gg}^{-1} M_g \\ Z_q M_q \end{bmatrix} \alpha + \begin{bmatrix} Z_n \\ 0 \end{bmatrix} \varepsilon_n + \begin{bmatrix} e_n \\ e_q \end{bmatrix}$$

With "Hybrid" Mixed Model Equations

$$\begin{bmatrix} X'X & X'ZM & X_n'Z_n \\ M'Z'X & M'Z'ZM + \phi & M_n'Z_n'Z_n \\ Z_n'X_n & Z_n'Z_nM_n & Z_n'Z_n + A^{nn} \lambda \end{bmatrix} \begin{bmatrix} b \\ \alpha \\ \varepsilon_n \end{bmatrix} = \begin{bmatrix} X'y \\ M'Z'y \\ Z_n'y_n \end{bmatrix}$$

$$where \ X = \begin{bmatrix} X_n \\ X_g \end{bmatrix}, Z = \begin{bmatrix} Z_n & 0 \\ 0 & Z_g \end{bmatrix}, M = \begin{bmatrix} M_n \\ M_g \end{bmatrix} = \begin{bmatrix} A_{ng}A_{gg}^{-1}M_g \\ M_g \end{bmatrix}, y = \begin{bmatrix} y_n \\ y_g \end{bmatrix}$$

with EBV given by

$$\widehat{u_g} = M_g \widehat{\alpha}$$

$$\widehat{u_g} = M_g \widehat{\alpha} + \widehat{\varepsilon}_g$$

NB Single-Step GBLUP
is a special case of the above
(but in this equivalent model no inversion is needed)

$$M_n = A_{ng}A_{gg}^{-1}M_g$$

If everyone is genotyped

$$\begin{bmatrix} X'X & X'ZM & X_n Z_n \\ M'Z'X & M'Z'ZM + \phi & M_n'Z_n'Z_n \\ Z_n'X_n & Z_n'Z_nM_n & Z_n'Z_n + A^{nn} \lambda \end{bmatrix} \begin{bmatrix} b \\ \alpha \\ \varepsilon_n \end{bmatrix} = \begin{bmatrix} X'y \\ M'Z'y \\ Z_n'y_n \end{bmatrix}$$

These are the MME that form the basis of BayesA, BayesB, BayesC etc.

If no one is genotyped

$$\begin{bmatrix}
X'X & X'ZM & X_n'Z_n \\
M'Z'X & M'Z'ZM + \phi & M_n'Z_n'Z_n \\
Z_n'X_n & Z_n'Z_nM_n & Z_n'Z_n + A^{nn}\lambda
\end{bmatrix}
\begin{bmatrix}
b \\
\alpha \\
\varepsilon_n
\end{bmatrix} =
\begin{bmatrix}
X'y \\
M'Z'y \\
Z_n'y_n
\end{bmatrix}$$

These MME form the basis of traditional pedigree-based BLUP

Invariant to Covariate Centering

Genotyped

$$egin{aligned} y_g &= \mathbf{1} \mu + X_g b + Z_g M_g lpha + e_g \ &= \mathbf{1} \mu + X_g b + Z_g \mathbf{1} c' lpha + Z_g (M_g - 1 c') lpha + e_g \ define \ t &= c' lpha \ y_g &= \mathbf{1} (\mu + t) + X_g b + Z_g (M_g - 1 c') lpha + e_g \ &= \mathbf{1} \mu^* + X_g b + Z_g M_g^c lpha + e_g \end{aligned}$$

.....when all animals genotyped (BayesA, BayesB etc)

But non-genotyped NOT invariant

 $egin{align*} Non-genotyped \ y_n &= \mathbf{1} \mu + X_n b + Z_n A_{ng} A_{gg}^{-1} M_g lpha + Z_n oldsymbol{arepsilon}_n + e_n \ &= \mathbf{1} \mu + X_n b + Z_n A_{ng} A_{gg}^{-1} \mathbf{1} c^{\dagger} lpha + Z_n A_{ng} A_{gg}^{-1} (M_g - 1 c^{\dagger}) lpha + Z_n oldsymbol{arepsilon}_n + e_n \ &= \mathbf{1} \mu + X_n b + Z_n A_{ng} A_{gg}^{-1} \mathbf{1} t + Z_n A_{ng} A_{gg}^{-1} M_g^c lpha + Z_n oldsymbol{arepsilon}_n + e_n \ \end{aligned}$

So combined analysis of genotyped and non-genotype animals need to include a covariate for t if there is arbitrary centering (unless t = 0)

Computational Aspects

- It is easy to compute $A_{ng}A_{gg}^{-1}M_g$
 - And this can be done in parallel
- The computing becomes easier (rather than more difficult or impossible) as more individuals are genotyped
- Readily caters for variable selection or mixture models (eg BayesB, BayesC)
- We believe this formulation is readily extended to multi-breed and multi-trait settings
- In an MCMC framework can provide PEV

Summary

- Genomic prediction is an immature technology
- Much effort is required to extend algorithms and to develop parallel computing procedures to implement the full range of multi-breed, multi-trait, maternal effects and other models that have been routinely applied to large-scale animal prediction in recent decades

Prediction of BVs

$$with \ EBV \ given \ by$$
 $\widehat{u_g} = M_g \widehat{lpha}$
 $\widehat{u_n} = M_n \widehat{lpha} + \widehat{oldsymbol{arepsilon}_n}$
 $or, with \ M_n = A_{ng} A_{gg}^{-1} M_g \widehat{lpha} + \widehat{oldsymbol{arepsilon}_n}$
 $= A_{ng} A_{gg}^{-1} \widehat{u_g} + \widehat{oldsymbol{arepsilon}_n}$